

# Econ 103: Introduction to Econometrics

## Lecture 17 — F-Tests & Joint Hypothesis Testing

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Reading: Hill, Griffiths & Lim (5th ed.), §6.1–6.2; Stock & Watson (4th ed.), §7.2.

## Where we are

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The  $t$ -test handles a **single** restriction — one “equals” sign, even one spanning several coefficients. But many questions are **joint**:

- Does advertising matter *at all* — is  $\beta_3 = 0$  and  $\beta_4 = 0$  in Big Andy’s quadratic model?
- Does a *group* of variables (socioeconomic controls, prices of substitutes) belong?
- Does the model explain *anything* — are **all** slopes zero?

These have **several** “equals” signs. A  $t$ -test can’t do them — and testing one at a time is unreliable. **Today: the  $F$ -test.**

- restricted vs. unrestricted models,
- overall significance and the  $t$ – $F$  relationship,
- testing economic restrictions (Cobb–Douglas constant returns to scale).

## Today's plan

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- ① Why a new test?
- ② The F-statistic: restricted vs. unrestricted
- ③ Overall significance and the t–F link
- ④ Testing economic restrictions

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Part 1

# **Why a new test?**

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## Joint hypotheses, and the one-at-a-time trap

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A **joint hypothesis** imposes  $J \geq 2$  restrictions at once, e.g.

$$H_0 : \beta_3 = 0 \text{ and } \beta_4 = 0 \quad \text{vs.} \quad H_1 : \beta_3 \neq 0 \text{ or } \beta_4 \neq 0.$$

### Why not just run two $t$ -tests?

- **Error rates compound.** Two 5%-tests don't give a 5% joint test — the chance of *some* false rejection is larger.
- **It misreads correlated regressors.** When two regressors are collinear, *each*  $t$  can be insignificant while the pair is jointly decisive. One-at-a-time would wrongly drop both.

We need a test that weighs *all* the restrictions together. That is the  $F$ -test.

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Part 2

# **The F-statistic: restricted vs. unrestricted**

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## Two models: with and without the restrictions

The  $F$ -test compares the fit of two nested models.

### Unrestricted (full)

$$\text{SALES} = \beta_1 + \beta_2 \text{PRICE} \\ + \beta_3 \text{ADVERT} + \beta_4 \text{ADVERT}^2 + e$$

Sum of squared errors:  $\text{SSE}_U$ .

### Restricted (impose $H_0$ )

Set  $\beta_3 = \beta_4 = 0$ :

$$\text{SALES} = \beta_1 + \beta_2 \text{PRICE} + e$$

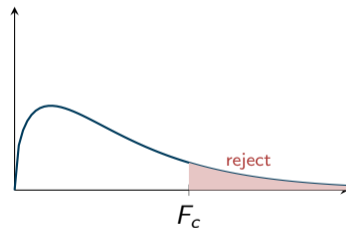
Sum of squared errors:  $\text{SSE}_R$ .

Dropping variables can only *worsen* fit, so  $\text{SSE}_R \geq \text{SSE}_U$  always. The question is whether the **increase** in SSE from imposing  $H_0$  is *large* (restrictions hurt  $\Rightarrow$  reject) or *small* (restrictions harmless  $\Rightarrow$  don't reject).

## The F-statistic

$$F = \frac{(SSE_R - SSE_U)/J}{SSE_U/(N - K)} \sim F_{(J, N-K)} \text{ under } H_0$$

- $J$  = number of restrictions (numerator df).
- $N - K$  = unrestricted df (denominator df).
- Numerator: **extra** error from imposing  $H_0$ , per restriction.
- Denominator: the model's own noise  $\hat{\sigma}^2$ .
- Large  $F \Rightarrow$  restrictions cost a lot  $\Rightarrow$  **reject**  $H_0$  if  $F \geq F_c$ .



F-distribution; reject in the right tail.

## Big Andy's: does advertising matter?

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Test  $H_0 : \beta_3 = 0, \beta_4 = 0$  (advertising, linear *and* quadratic, irrelevant) against “at least one nonzero.” Here  $J = 2, N = 75, K = 4$ :

$$\text{SSE}_U = 1532.08, \quad \text{SSE}_R = 1896.39.$$

$$F = \frac{(1896.39 - 1532.08)/2}{1532.08/(75 - 4)} = 8.44.$$

The 5% critical value is  $F_{(0.95,2,71)} = 3.13$ ;  $p = 0.0005$ .

Since  $8.44 > 3.13$ , **reject  $H_0$**  — advertising does affect sales. (We couldn't have learned this cleanly from the two separate  $t$ 's, since ADVERT and ADVERT<sup>2</sup> are collinear.)

Equivalent  $R^2$  form (S&W):  $F = \frac{(R_U^2 - R_R^2)/J}{(1 - R_U^2)/(N - K)}$  — same number, computed from fit instead of SSE.

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Part 3

## **Overall significance and the t–F link**

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## Testing the whole model at once

The most-reported  $F$ -test asks: do the regressors *jointly* explain anything?

$$H_0 : \beta_2 = \beta_3 = \cdots = \beta_K = 0 \quad (\text{the model is worthless}).$$

The restricted model is just  $y_i = \beta_1 + e_i$ , fitted by  $\bar{y}$ , so  $SSE_R = SST$ . The statistic specializes to

$$F = \frac{(SST - SSE)/(K - 1)}{SSE/(N - K)} \sim F_{(K-1, N-K)}.$$

### Example (Big Andy's overall $F$ )

With  $SST = 3115.48$ ,  $SSE = 1532.08$ ,  $K = 4$ :

$$F = \frac{(3115.48 - 1532.08)/3}{1532.08/71} = 24.46 \gg F_c = 2.73.$$

*Reject* — at least one of  $PRICE$ ,  $ADVERT$ ,  $ADVERT^2$  matters. This is the **overall significance**  $F$  printed on every regression.

## $t$ and $F$ : when are they the same?

For a single restriction ( $J = 1$ ): identical

A two-tailed  $t$ -test and the  $F$ -test give the **same** answer, because

$$F = t^2 \quad \text{and} \quad F_c = t_c^2.$$

(Big Andy,  $H_0 : \beta_2 = 0$ :  $t = -7.30$ ,  $t^2 = 53.4 = F$ .) Same  $p$ -value, same conclusion.

But two crucial places where *only* the right tool works:

- **One-tailed** tests ( $H_1 : \beta > c$ ): use  $t$ . Since  $F = t^2$  squares away the sign,  $F$  *cannot* do one-sided.
- **Joint** tests ( $J \geq 2$ ): use  $F$ . There is no single  $t$ .

Custom: test single restrictions with  $t$ , joint restrictions with  $F$ .

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Part 4

# Testing economic restrictions

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## Beyond “= 0”: restrictions from theory

The real power of the  $F$ -test: the restrictions can be *any* linear equalities theory hands us — not just “this coefficient is zero.”

### Cobb–Douglas and constant returns to scale

A production function  $Q = AL^{\beta_2}K^{\beta_3}$  in logs is

$$\ln Q = \beta_1 + \beta_2 \ln L + \beta_3 \ln K + e.$$

**Constant returns to scale** — doubling all inputs doubles output — is exactly

$$H_0 : \beta_2 + \beta_3 = 1.$$

Impose it (a restricted model with one fewer free parameter), get  $SSE_R$ , and form the  $F$  with  $J = 1$ . Increasing returns if the data reject in favor of  $\beta_2 + \beta_3 > 1$ .

## More restriction examples

### No money illusion (HGL beer demand)

Log-log demand  $\ln Q =$

$$\beta_1 + \beta_2 \ln P_B + \beta_3 \ln P_L + \beta_4 \ln P_R + \beta_5 \ln I + e.$$

Scaling all prices *and* income shouldn't change quantity, so

$$H_0 : \beta_2 + \beta_3 + \beta_4 + \beta_5 = 0.$$

### Is \$1,900 the optimal ad spend?

Andy's optimum satisfies

$$\beta_3 + 2\beta_4 \text{ADVERT} = 1. \text{ At ADVERT} = 1.9,$$

$$H_0 : \beta_3 + 3.8\beta_4 = 1.$$

$F = 0.94 < 3.98$  — fail to reject; \$1,900 is compatible with the data.

### How it's done in practice

Either rewrite the model to embed the restriction and re-estimate (get  $SSE_R$ ), or just hand the restriction to software — it computes the  $F$  (a **Wald test**) and  $p$ -value directly.

## Joint restrictions can mix several conjectures

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Nothing stops  $H_0$  from bundling *different* economic claims. Andy plans staffing on two assumptions at once: \$1,900 is optimal *and* sales at PRICE = 6, ADVERT = 1.9 average \$80,000:

$$H_0 : \beta_3 + 3.8\beta_4 = 1 \quad \text{and} \quad \beta_1 + 6\beta_2 + 1.9\beta_3 + 3.61\beta_4 = 80.$$

Two restrictions ( $J = 2$ )  $\Rightarrow$  an  $F$ -test (no  $t$  possible). Here  $F = 5.74$ ,  $p = 0.005$  — **reject**: the two plans are *jointly* incompatible with the data, even if each alone might survive.

This is the everyday use of  $F$ -tests in research: bundling a model's theoretical restrictions and asking whether the data can live with all of them together.

## Recap

### The $F$ -test

- joint  $H_0$  ( $J \geq 2$  restrictions)  $\Rightarrow$  need  $F$ , not many  $t$ 's
- $F = \frac{(\text{SSE}_R - \text{SSE}_U)/J}{\text{SSE}_U/(N - K)} \sim F_{(J, N-K)}$
- big SSE jump from  $H_0 \Rightarrow$  reject
- Andy ads:  $F = 8.44$ , reject

### Overall significance

- all slopes = 0; restricted model =  $\bar{y}$
- Andy:  $F = 24.46$ , model significant

### $t$ vs. $F$

- $J = 1$ :  $F = t^2$ , equivalent (two-tail)
- one-tail  $\rightarrow t$  only; joint  $\rightarrow F$  only

### Economic restrictions

- CRS:  $\beta_2 + \beta_3 = 1$
- no money illusion:  $\sum \beta = 0$
- optimal ad:  $\beta_3 + 3.8\beta_4 = 1$  (not rejected)

## Next time (Lecture 18): which variables belong?

$F$ -tests assumed a model. But choosing it is the hard part: [omitted-variable bias](#) vs. [irrelevant variables](#), and tools to decide — adjusted  $R^2$ , AIC/BIC, the RESET test, and residual diagnostics.

Questions?